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Global genetic diversity and historical demography of the Bull Shark

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Abstract

Aim: Biogeographic boundaries and genetic structuring have important effects on the inferences and interpretation of effective population size (N_e) temporal variations, a key genetics parameter. We reconstructed the historical demography and divergence history of a vulnerable coastal high-trophic shark using population genomics and assessed our ability to detect recent bottleneck events.

Location: Western and Central Indo-Pacific (IPA), Western Tropical Atlantic (WTA) and Eastern Tropical Pacific (EPA).

Taxon: *Carcharhinus leucas* (Müller & Henle, 1839).

Methods: A DArTcap™ approach was used to sequence 475 samples and assess global genetic structuring. Three demographic models were tested on each population, using an ABC-RF framework coupled with coalescent simulations, to investigate within-cluster structure. Divergence times between clusters were computed, testing multiple scenarios, with *fastsimcoal*. N_e temporal variations were reconstructed with STAIRWAYPLOT. Coalescent simulations were performed to determine the detectability of recent bottleneck under the estimated historical trend for datasets of this size.

Results: Three genetic clusters corresponding to the IPA, WTA and EPA regions were identified, agreeing with previous studies. The IPA presented the highest genetic diversity and was consistently identified as the oldest. No significant within-cluster structuring was detected. N_e increased globally, with an earlier onset in the IPA, during the last glacial period. Coalescent simulations showed that weak and recent bottlenecks could not be detected with our dataset, while old and/or strong bottlenecks would erase the observed ancestral expansion.

Main Conclusions: This study further confirms the role of marine biogeographic breaks in shaping the genetic history of large mobile marine predators. N_e historical increases in N_e are potentially linked to extended coastal habitat availability. The limited within-cluster population structuring suggests that N_e can be monitored over ocean basins. Due to insufficient amount of available genetic data, it cannot be concluded whether overfishing is impacting Bull Shark genetic diversity, calling for whole-genome sequencing.

KEYWORDS

Carcharhinidae, coalescent simulations, DArTcap, demographic history, marine biogeography

1 | INTRODUCTION

Three biogeographic boundaries have been promoting speciation in the marine realm since the early Neogene: the Eastern Tropical Pacific open ocean, the Benguela Current and the Isthmus of Panama (Cowman & Bellwood, 2013; Lessios, 2008; O'Dea et al., 2016; Waters, 2008). Their gradual formation has separated a once continuous tropical ocean linked through the Tethyan Seaway and several seaways, connecting current Western Tropical Atlantic (WTA) to the Western and Central Pacific (IPA) between the Triassic and the Pliocene (Hou & Li, 2018; Popov et al., 2004). On the eastern side of this ocean, the Eastern Pacific open ocean has been preventing eastward species colonization from the IPA to the Eastern Tropical Pacific (EPA) for at least 65 million years before present (B.P.; Cowman & Bellwood, 2013). On its western side, the closure of the Tethyan Seaway at the end of the Middle Miocene (Sun et al., 2021) and the formation of the Benguela Current during the Pliocene (Jung et al., 2014) isolated the IPA from WTA. Finally, the formation of the Isthmus of Panama at the end of the Pliocene isolated the EPA from the WTA (O'Dea et al., 2016). These boundaries limited or stopped gene flow between populations, impacting genetic structure and diversity. Known biogeographic breaks not only provide foundation to identify biodiversity patterns but also help in delineating conservation regions when studying their effects on species connectivity (Fredston-Hermann et al., 2018; Norris, 2004).

Effective species' conservation and management require understanding of their population dynamics, biogeographical ranges, life history traits and genetic connectivity (Green et al., 2014; Hohenlohe et al., 2021; Young et al., 2006). These factors shape temporal variations in population census and effective sizes (N_e). Census size is a parameter difficult to measure for vagile or rare species (Gerber et al., 2014) and it does not inform on adaptive potential (Reed & Frankham, 2016). Conversely, N_e and its temporal variation can be estimated using molecular markers. This ultimately enables an understanding of environmental or human-induced factors influencing such variation, providing clues on adaptive potential and species management planning (Luikart et al., 2010; Nadachowska-Brzyska et al., 2021; Ouborg et al., 2010). However, N_e estimations require knowledge of the demographic history, as spatial structure influences N_e and may bias inferences (Arredondo et al., 2021;

Chikhi et al., 2010; Lesturgie, Lainé, et al., 2022; Lesturgie, Planes, & Mona, 2022; Maisano Delser et al., 2019; Mazet et al., 2016). Previous studies showed through theoretical and simulation arguments that incorrect modelling of population structure may lead to inaccurate historical demography interpretation (Chikhi et al., 2010; Lesturgie, Lainé, et al., 2022; Lesturgie, Planes, & Mona, 2022; Maisano Delser et al., 2019; Wakeley, 2009).

Elasmobranchs are among the most threatened marine organisms (Dulvy et al., 2021). Many species exhibit late maturity, low fecundity, long gestation and slow growth, making them susceptible to overfishing (Adams, 1980; Cortés, 2000). Moreover, the common reliance on nursery areas (Heithaus, 2005) and philopatric behaviour (Chapman et al., 2015) increase the risk of local extinctions. Many modern elasmobranch groups predate the Tethyan closure, with this subclass probably already widely distributed by the Lower Jurassic (Maisey, 2012). Biogeographic barriers have different effects on elasmobranch populations, mainly due to their reproductive ecology and physiology (Kottillil et al., 2023). While the Benguela Current is a strong barrier for many organisms (Teske et al., 2011), partial migration from the IPA to the WTA has already been documented in some sharks (Lesturgie, Lainé, et al., 2022; Lesturgie, Planes, & Mona, 2022). Strong barriers such as the Isthmus of Panama or the Eastern Pacific open ocean have promoted genetic differentiation and even speciation of shark populations (Gonzalez et al., 2021; Pazmiño et al., 2018). Based on mitochondrial DNA (mtDNA) data, coastal or demersal species tend to present genetic structuring at small geographic scales (Hirschfeld et al., 2021; Momigliano et al., 2017; Vignaud et al., 2014) while pelagic or semi-pelagic species show low structuring between and within ocean basins (Bailleul et al., 2018; Pirog, Jaquemmet, et al., 2019). Until recently, most elasmobranch genetic studies relied on traditional markers, that is, mtDNA and microsatellites (Phillips et al., 2021). These markers represent small portions of a genome, allowing only partial reconstructions of a species' evolutionary history. Nowadays, the popularity of genotyping-by-sequencing (GBS) approaches has fuelled genomic studies in non-model organisms (e.g. Combosch & Vollmer, 2015; Harvey & Brumfield, 2015). However, few elasmobranchs have benefited yet (see Devloo-Delva et al., 2023; Feutry et al., 2020; Glaus et al., 2020; Lesturgie et al., 2023; Lesturgie, Lainé, et al., 2022; Lesturgie, Planes, & Mona, 2022; Maisano Delser et al., 2016, 2019;

Pazmiño et al., 2018). Filling this gap will address several evolutionary questions and prompt refined conservation management planning.

The Bull Shark *Carcharhinus leucas* (Müller & Henle, 1839) is an euryhaline, globally distributed, migratory species inhabiting tropical to warm temperate waters (Compagno, 1990). The earliest fossils of this species date 23 million B.P. and are present across what was the Tethys Sea, from Peru to the Mekong River (Gausmann, 2021). This species can travel along continental coasts (Espinoza et al., 2016, 2021; Heupel et al., 2015), into freshwater rivers (Werry et al., 2012) and across open ocean (Lea et al., 2015). Its trophic position in food webs, combined with its movement, makes the species ecologically important. Females rely on coastal nurseries (Sandoval Laurraquiao-Alvarado et al., 2019; Tillett et al., 2012) and some studies hypothesized a tendency for philopatry, based on telemetry and genetic data (Espinoza et al., 2016; Pirog, Jaquemmet, et al., 2019; Pirog, Ravigné, et al., 2019). The Bull Shark evolutionary history has been investigated using traditional molecular markers (Karl et al., 2011; Pirog, Ravigné, et al., 2019; Sandoval Laurraquiao-Alvarado et al., 2019; Testerman, 2014) and GBS data (Devloo-Delva et al., 2023; Glaus et al., 2020), but a detailed modelling of its N_e historical trajectory and the timing of divergence between inferred genetic clusters is lacking. Moreover, N_e temporal trends and estimates are inconsistent due to the limits and variety of molecular markers used to date (Karl et al., 2011; Pirog, Ravigné, et al., 2019; Sandoval Laurraquiao-Alvarado et al., 2019; Testerman, 2014). An assessment of current demographic trends is crucial, as populations have declined in the IPA (based on catch data); *Carcharhinus leucas* was recently assessed as Vulnerable by the International Union for the Conservation of Nature (IUCN) Red List of Threatened Species (Rigby et al., 2021). This decline is probably due to overfishing, as it is among the most traded species (Cardeñosa et al., 2018, 2022; Cardeñosa, Fields, et al., 2020; Cardeñosa, Shea, et al., 2020; Fields et al., 2018).

The present study aims to: (1) identify the most likely evolutionary divergence scenario that may have shaped the observed genetic structure of *C. leucas*; (2) reconstruct the historical variation of N_e in the identified clusters; and (3) test whether recent bottlenecks could be detected given the observed genetic diversity and sample sizes in this study. Results will inform management and conservation actions by providing a first estimate of *C. leucas* N_e and its historical trend on a global scale while assessing our ability to monitor N_e with the available genomic data.

2 | MATERIALS AND METHODS

2.1 | Sample collection and DNA extraction

A subsample of the dataset of Devloo-Delva et al. (2023) was used for this study, representing 475 *C. leucas* sampled between 1985 and 2019 from 18 locations covering its distribution (except for West Africa; Data S1). DNA was extracted with the Qiagen Blood and Tissue kit, following standard protocol (Qiagen Inc., Valencia, California, USA). After bait design and bioinformatic filtering (see following sections), the dataset comprised 16 sampling locations with at least five individuals (309 individuals; Figure 1, Table 1) covering the WTA, IPA and EPA. Sampling locations with mostly adults were preferentially selected to limit relatedness effects.

2.2 | SNP selection for bait design

The approach used for bait design is described in Devloo-Delva et al. (2023). Briefly, a subset of 219 sample libraries were genotyped using the DArTseq™ approach (Cruz et al., 2013; Feutry et al., 2017, 2020, Data S1). From this dataset, 3400 loci of 70bp were randomly

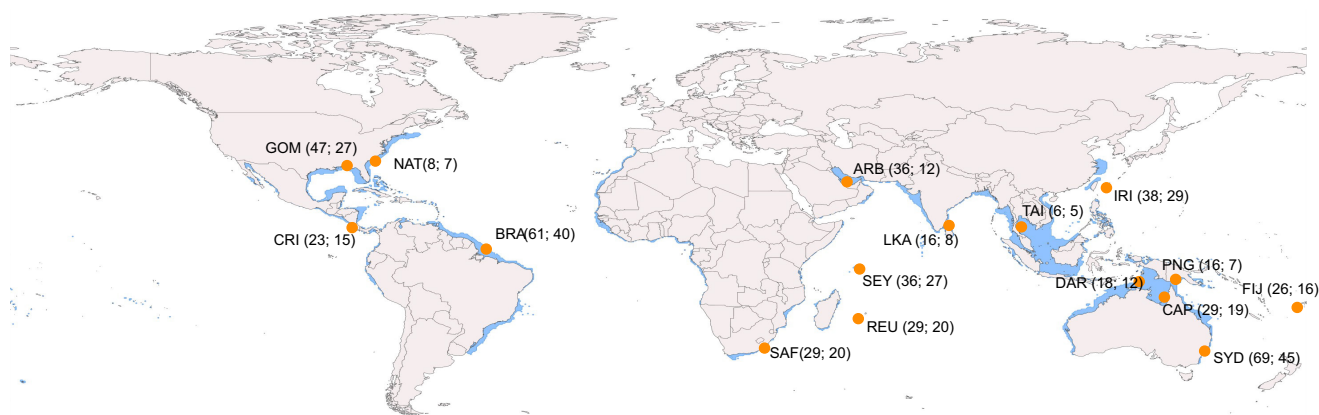


FIGURE 1 Distribution range (blue) and sampling locations (orange) of *Carcharhinus leucas* populations. In parentheses, the number of individuals sequenced (left) and the number of individuals that passed bioinformatic filtering (right). ARB, Arabian/Persian Gulf and Arabian Sea; BRA, Brazil; CAP, Cap York; CRI, Costa Rica; DAR, Darwin; FIJ, Fiji; GOM, Gulf of Mexico; IRI, Iriomote Island; LKA, Sri Lanka; NAT, U.S. Atlantic coast; PNG, Papua New Guinea; REU, Reunion Island; SAF, South Africa; SEY, Seychelles; SYD, Sydney; TAI, Thailand. Distribution range is based on the IUCN SSC Shark Specialist Group 2020. *Carcharhinus leucas*. The IUCN Red List of Threatened Species. Version 2022-2. <https://www.iucnredlist.org>. Downloaded on 20 May 2023.

TABLE 1 Summary statistics for each *Carcharhinus leucas* population-specific dataset used in this study and ABC-RF estimation of the best demographic model. We only reported the estimated parameter values of datasets showing a posterior probability of ≥ 0.5 . Prior distributions are set to uniform for all estimated parameters.

Sampling region	Sampling site	Year of sampling	Code	N_{seq}	N_{bio}	N_{loc}	N_{SNP}	Mean coverage	TD	θ_{π}	θ_w	Model (prob.)	N_{mod} [95% Conf. Int.]	T_{col} [95% Conf. Int.]	N_{anc} [95% Conf. Int.]
EPA	Costa Rica	2017–2018	CRI	17	15	5145	695	24.499	-0.75	3.93×10^{-4}	4.87×10^{-4}	SST (0.373)			
WTA	Gulf of Mexico	2011–2017	GOM	47	27	2981	549	27.824	-1.02	4.14×10^{-4}	5.77×10^{-4}	NS (0.973)	16,040 [5130–37,892]	155,535 [4216–1,792,297]	4717 [226–37,537]
WTA	Brazil	2003–2005	BRA	61	40	3408	793	26.347	-1.26	4.28×10^{-4}	6.71×10^{-4}	NS (0.436)			
WTA	U.S. Atlantic coast	1987–2015	NAT	8	7	3569	417	26.954	-0.54	4.63×10^{-4}	5.25×10^{-4}	SST (0.337)			
IPA	South Africa ^a	2009–2015	SAF	28	20	3106	909	28.130	-0.77	7.82×10^{-4}	9.83×10^{-4}	NS (0.594)	17,719 [12,410–39,175]	125,536 [18,558–389,778]	5334 [170–8395]
IPA	Arabian/Persian Gulf and Arabian Sea ^a	2010–2012	ARB	23	12	5312	1108	25.114	-0.66	6.68×10^{-4}	7.98×10^{-4}	NS (0.583)	17,454 [8403–37,213]	99,533 [12,985–304,777]	5043 [482.33–7223]
IPA	Seychelles ^a	2013–2016	SEY	36	27	3699	1167	26.266	-1.03	7.09×10^{-4}	9.89×10^{-4}	NS (0.619)	21,084 [12,472–38,888]	110,483 [15,963–304,781]	4545 [894–7565]
IPA	Reunion Island ^a	2013–2017	REU	28	20	3095	876	28.084	-0.80	7.49×10^{-4}	9.51×10^{-4}	NS (0.666)	17,071 [12,123–38,887]	130,408 [23,300–332,737]	5139 [966–8395]
IPA	Sri Lanka ^a	2017–2018	LKA	12	8	4363	875	26.916	-0.57	7.52×10^{-4}	8.63×10^{-4}	NS (0.502)	22,160.02 [12,152–47,651]	107,423 [16,294–263,978]	5738 [1822–8044]
IPA	Thailand	Not recorded	TAI	6	5	4242	617	26.045	-0.61	6.45×10^{-4}	7.34×10^{-4}	NS (0.124)			
IPA	Darwin	2008	DAR	17	12	4482	949	26.884	-0.60	6.90×10^{-4}	8.10×10^{-4}	NS (0.481)			
IPA	Papua New Guinea	2018–2019	PNG	12	7	6799	1072	25.470	-0.53	6.26×10^{-4}	7.08×10^{-4}	NS (0.471)			
IPA	Cape York ^a	2002–2009	CAP	27	19	3806	1081	27.302	-0.90	7.35×10^{-4}	9.66×10^{-4}	NS (0.638)	19,887 [12,122–39,175]	110,902 [18,568–304,774]	5554 [1332–7780]
IPA	Sydney ^a	2011–2019	SYD	69	45	3138	1127	28.121	-0.93	7.38×10^{-4}	1.01×10^{-3}	NS (0.620)	16,367 [12,152–29,654]	107,193 [18,668–304,773]	4912 [1192–8216]
IPA	Iriomote Island	2014–2016	IRI	38	29	4210	897	26.668	-0.30	6.02×10^{-4}	6.58×10^{-4}	NS (0.658)	6727 [5823–9026]	1,603,507 [69,952–3,824,732]	18,639 [226–46,643]
IPA	Fiji	2016–2017	FIJ	25	16	2583	558	28.003	-0.40	6.87×10^{-4}	7.66×10^{-4}	NS (0.562)	11,848 [6854–42,923]	864,559 [12,595–3,730,256]	9451 [279–44,150]
Priors													0.5–50,000	13–3,900,000	0.5–50,000

Abbreviations: EPA, Eastern Tropical Pacific; IPA, Western and Central Indo-Pacific; N_{anc} , ancestral N_e expressed in diploid genotypes; N_{bio} , number of individuals that passed bioinformatic filtering; N_{loc} , number of loci retained after bioinformatic filtering; N_{mod} , modern N_e expressed in diploid genotypes; N_{seq} , number of individuals sequenced; N_{SNP} , number of SNPs retained after bioinformatic filtering; T_{col} , time of the instantaneous change from N_{mod} to N_{anc} expressed in years before present; TD, Tajima's D; WTA, Western Tropical Atlantic; θ_w , Watterson's theta; and θ_{π} , mean pairwise difference.

^aPopulation used to design the simulation study.

selected for DNA capture bait development. The DArTcap™-enriched libraries were sequenced on an Illumina HiSeq 2500.

2.3 | Bioinformatics

Reads were demultiplexed with DArTsoft14™ and analysed using STACKS 2.5 (Rochette et al., 2019). STACKS clustering parameters were optimized as recommended by Paris et al. (2017). First, the de novo pipeline was run on a randomly selected sampling site (Seychelles) with different combinations of m (minimum number of raw reads to form a stack; from 3 to 10), M (number of mismatches between stacks within an individual to merge stacks; set to 4, 6 or 8) and n (number of mismatches between stacks in different individuals; equal to M). The number of polymorphic loci, SNPs, the nucleotide diversity θ_π and θ_w (Watterson, 1975) were compared between parameter combinations, allowing up to 20% of missing data per locus. We selected the parameters $m=3$, $M=4$ and $n=4$ which maximized the number of loci retrieved without over-splitting the dataset. Using these parameters, the de novo pipeline was run on individuals belonging to sampling locations with more than five individuals. Loci were first filtered using the *population* function to discard: (i) SNPs with heterozygosity rate >0.8 ; and (ii) SNPs with more than 20% missing data in any sampling site. Finally, we filtered the dataset with a custom R script (R Core Team, 2022) to discard: (i) loci with more than five SNPs (after checking the empirical distribution of SNPs per locus); (ii) SNPs with average coverage $<10\times$ or $>60\times$ (after checking the empirical distribution); and (iii) individuals with more than 10% missing data.

Additional filters were applied depending on downstream analyses. To analyse population genetic structure, one random SNP per locus was retained (to avoid linkage disequilibrium) with a minor allele frequency >0.05 (hereafter, the *global* dataset). To estimate the genetic diversity in each population and model historical demography, all SNPs without missing data were retained (called *population* dataset).

2.4 | Population structure

As this study uses a subset of an existing dataset (Devloo-Delva et al., 2023), standard population structure analyses were performed to assess the concordance and robustness of previous results.

The global genetic structure was evaluated using a hierarchical approach in fastSTRUCTURE 1.0 (Raj et al., 2014). For the *global* dataset and each sub-dataset, three independent runs were performed with K varying between 1 and 10. This was performed until no sub-clustering was detected (i.e. optimal $K=1$). The expected admixture proportions inferred by fastSTRUCTURE were visualized with DISTRUCT (Rosenberg, 2004).

Population clustering was further investigated using a discriminant analysis of principal component (DAPC) with the 'adegenet' R package (Jombart, 2008). The decrease in Bayesian

information criterion (BIC) values was examined to identify the optimal K (Jombart et al., 2010). The *dapc* function was executed using the chosen K value, retaining the axes of PCA explaining $\geq 80\%$ of the total variance. Pairwise fixation indices (F_{ST} ; Reynolds et al., 1983) between sampling locations were calculated using the 'diveRsim' R package (Keenan et al., 2013), with significance tested after 1000 permutations.

2.5 | Demographic inferences

Historical demography was explored using the STAIRWAYPLOT 2.0 (Liu & Fu, 2020). The STAIRWAYPLOT models the folded site frequency spectrum (SFS) to infer coalescence rate changes through time. If individuals come from a panmictic population, the coalescence rate can be converted to N_e using a generation time and a mutation rate. We applied a generation time of 13 years, computed as the average age of sexual maturity of 15 years in the Atlantic (Branstetter & Stiles, 1987) and ~12 years in the Indian Ocean (Hoarau et al., 2021). It should be noted that this arbitrary value represents the minimal age at which an individual could contribute to the genetic diversity of the next generation [the IUCN reports a generation time of 22.7 years (Rigby et al., 2021)]. We applied the mutation rate estimated by Lesturgie, Lainé, et al. (2022); Lesturgie, Planes, & Mona, (2022) based on *Carcharhinus melanopterus* RAD-seq data after scaling it to account for the generation time of *C. leucas*. This resulted in a mutation rate of 2.509×10^{-8} per site per generation.

Population structure can bias the estimation of temporal N_e variation based on models assuming panmixia (Chikhi et al., 2010; Lesturgie, Lainé, et al., 2022; Lesturgie, Planes, & Mona, 2022; Maisano Delser et al., 2019; Wakeley, 2009). It is therefore important to test for population structuring before interpreting the reconstructed N_e . The approach proposed by Lesturgie, Lainé (2022); Lesturgie, Planes, & Mona, (2022) was used in addition to the clustering analyses. We devised three demographic models to test if the summary statistics observed in each sampling location (deme) are more likely to be described by an unstructured model (i.e. a panmictic population) or a meta-population represented by an array of demes exchanging migrants either under a finite island or a steppingstone model. Each deme was analysed separately, and the probability of each of the three models is evaluated using an approximate Bayesian computation (ABC) framework. This approach has been shown to capture major features of the gene genealogy of a sample of lineages, that is, if they originate from a single panmictic deme or from a deme belonging to a meta-population (Maisano Delser et al., 2019; Peter et al., 2010).

The non-structured model (NS) represents a modern population of constant size N_{mod} switching instantly to an ancestral population of size N_{anc} at T_c generations in the past. The finite island model (FIM) represents an array of 100 demes each with constant size N_{mod} and exchanging N_{mig} migrants per generation (lineages were sampled from one random deme of the array); backward in time, all demes merged

instantaneously at T_c generations ago into a single population of size N_{anc} . The stepping-stone model (SST) resembles FIM, the difference being that populations only exchange migrants with their four closest neighbours and the lineages were sampled from one of the central demes of the array (Data S2). A total of 50,000 coalescent simulations for each model were performed with *fastsimcoal* 2.7.05 (Excoffier et al., 2013), extracting parameters from prior distributions (Table 1). Each sampling site was analysed separately, and simulations reproduced the exact number of individuals and loci observed in its corresponding *population* dataset. Model election and parameter estimations were based on the following set of summary statistics: the folded SFS, θ_π , θ_w (Tajima, 1989), and the number of segregating sites (S). Summary statistics were considered for both model selection and parameter estimation. Model selection was performed using the random forest (RF) classification approach implemented in the 'abcRF' R package (Pudlo et al., 2016). RF was trained using the simulated datasets, represented by the vector of summary statistics. The observed data were then assigned to one of the three model. We considered the model assignment reliable if its probability was >0.5 . The demographic parameters of the best model were then estimated with the 'abcRF' regression method (Raynal et al., 2019). The number of trees of each RF algorithm was chosen by monitoring the out-of-bag error (Pudlo et al., 2016). A confusion matrix was also generated during the model selection procedure to determine its performance: simulated datasets were assigned to one of the three models under investigation following the same procedure applied to the observed data. This allows to test the robustness of our procedure within the space of prior parameters chosen.

2.6 | Simulation study—Detection of recent bottleneck

The detectability of recent bottlenecks (5 to 1500 generations) was explored by running the STAIRWAYPLOT on simulated datasets having a number of individuals and loci consistent with the *population* datasets (see below). We focused on recent bottlenecks in populations experiencing a demographic history consistent with the one reconstructed here. According to our results, the demographic trajectories for most sampling locations could be described by a NS model with an ancestral N_e of 5000 individuals switching 6000 generations ago to a modern N_e of 16,000 individuals. These values were based on averages taken from both the STAIRWAYPLOT and ABC-RF results at sampling locations for which the NS model had a posterior probability >0.5 . *fastsimcoal* was used to run coalescent simulations under this NS model to which an instantaneous bottleneck was added (hereafter, NS_{BOT} model). Two hundred and four scenarios were investigated (Table 3, Data S3), combining variations in: (i) number of sampled individuals (5, 10, 15 or 20); (ii) number of sampled independent loci (1000, 5000 or 10,000 loci of 100bp); (iii) onset of the bottleneck (called T_{BOT}) in number of generations ago, taking values of 0, 5 [65 B.P., the beginning of industrial fishing (Mansfield, 2010)], 50 (650 B.P., an intermediate value within the last

millennium), 450 [5850 B.P., the end of the Holocene climate optimum (Summerhayes & Charman, 2015)] and 1500 [19,500 B.P., the end of the Last Glacial Maximum (LGM; Clark et al., 2009)]; and (iv) strength of the bottleneck (called BOT) which was set to decrease modern N_e 0, 5, 10, 50 or 100 times. Ten simulations were replicated per parameters combination, and the same summary statistics as in the real data (θ_π , θ_w , TD and S) were computed. STAIRWAYPLOT was then run on all replicates and the average of the estimated values was plotted.

2.7 | Ancestral divergence

fastsimcoal was used to investigate the timing of divergence among the three main biogeographic regions (i.e. WTA, EPA and IPA), corresponding to the three genetic clusters identified (see results and Data S6), in agreement with Devloo-Delva et al. (2023). This method uses a composite likelihood approach to optimize population demographic parameters under a defined scenario. The likelihood is computed by comparing the observed SFS to the one expected given a specific combination of demographic parameters, which is obtained by means of coalescent simulations. To maximize the number of loci without missing data and obtain a balanced sampling for each region, 15 individuals were randomly sampled from each genetic cluster (the U.S. Atlantic coast population was excluded to use individuals sampled on the same timeframe), hereafter the *divergence* dataset. We used identical filters as for other historical demographic analyses and we calculated the pairwise folded two-dimensional SFS (2D-SFS). Twenty-two scenarios were tested (Data S4). First, we tested the most likely population tree topology, that is, a synchronous or a sequential divergence between the genetic clusters (Figure 2). Then, we tested for continuous (symmetrical or asymmetrical) gene flow among clusters for the best population tree topology. We further tested the likelihood of a secondary contact between EPA and WTA after a complete isolation, potentially initiated by the opening of the Panama Canal around eight generations ago. The secondary contact model was tested within all tree topologies.

The likelihood of each scenario and its parameter values was assessed after selecting the best of 100 independent runs. The likelihood was evaluated by 250,000 simulations for each parameter combination and maximized by implementing 50 expectation-conditional-maximization cycles (Meng & Rubin, 1993). The range of modern and ancestral N_e (N_e and N_{anc} , Figure 2) was bounded between 50 and 50,000 individuals for each cluster. Divergence times (T_x , Figure 2) ranged between 100 and 100,000 generations (1300 to 1300,000 B.P.). Per generation migration rates were investigated between 10^{-7} and 0.01. *fastsimcoal* can explore values beyond boundaries if the likelihood increases. The run with the highest likelihood within each scenario was extracted to perform model selection using the Akaike information criterion (AIC). Parameters' confidence intervals were calculated with a parametric bootstrap approach: 100 datasets were simulated using the maximum-likelihood values of the best scenario, and then the 2D-SFS was computed for each

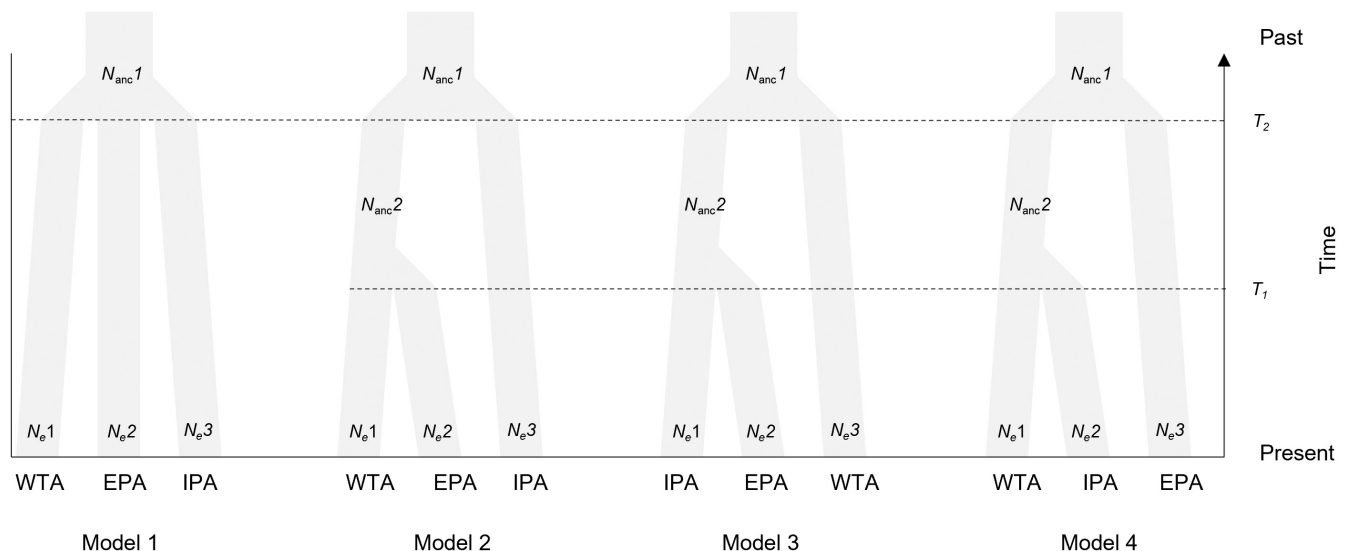


FIGURE 2 Population tree topologies of the four scenarios relating to the three genetic clusters identified in this study. N_e , modern effective population size, N_{anc} , ancestral effective population size, T_1 , first time of divergence; T_2 , second time of divergence; EPA, Eastern Tropical Pacific; IPA, Western and Central Indo-Pacific; WTA, Western Tropical Atlantic.

pair of population comparisons. Finally, *fastsimcoal* was run on each of these replicates with the same condition as for the real data. The best run out of 100 for each replicate was chosen to build the final confidence interval.

3 | RESULTS

3.1 | Genotyping of DArTcap data and datasets

After filtering, 734 polymorphic loci were recovered from the *global* dataset, with a mean read depth of $\sim 37\times$ ($SE=0.11$) per locus. Between 558 and 1167 (mean $\pm SE=938.82 \pm 61.24$), SNPs per sampling site were obtained in the *population* datasets (Table 1), with a mean read depth per locus per sampling site ranging from $24.50\times$ ($SE=0.15$) in Costa Rica to $28.13\times$ ($SE=0.14$) in South Africa. Finally, 715 polymorphic loci were obtained in the *divergence* dataset, with a mean depth of $25.81\times$ ($SE=0.29$) per locus.

3.2 | Population clustering and genetic connectivity

Strong genetic differentiation was identified among IPA, WTA and EPA. *fastSTRUCTURE* analyses suggested $K=2$ as the best number of clusters for the *global* dataset (Data S7a): WTA and EPA individuals clustered together and the IPA formed a second cluster. When analysing only WTA and EPA, *fastSTRUCTURE* identified two genetic clusters, matching the individuals' biogeographic origin. No further sub-clustering was detected within these regions. EPA was not run alone as it consists of a single sampling location. DAPC did not identify a single best solution according to the BIC, but its visual

inspection suggested K equal to 2 or 3 as the most likely values (Data S7b), consistent with *fastSTRUCTURE*. For $K=2$, DAPC identified one cluster with only IPA individuals and the other one with WTA+EPA individuals. For $K=3$, the first axis separated the WTA and EPA individuals from the IPA, while the second axis segregated individuals from the EPA, explaining $>95\%$ of the total variance.

The analysis of molecular variance computed using the three biogeographic regions as groups (in agreement with the clustering results) revealed that 54.81% of the total variance is partitioned in the between-region component ($p<0.005$), compared to 1.01% ($p<0.005$) in the between-sampling locations within regions component. The remaining genetic variation was found within sampling locations (45.2%, $p<0.005$).

All pairwise differentiation tests between sampling locations from different biogeographic regions showed significant F_{ST} values (range: 0.33–0.69). The mean genetic differentiation between sampling locations from the EPA and WTA (mean $F_{ST}=0.36$, range: 0.33–0.39) was lower than the mean differentiation between sampling locations from the IPA and the other two biogeographic regions (IPA vs. EPA mean $F_{ST}=0.62$, range: 0.56–0.69; IPA vs. WTA mean $F_{ST}=0.61$, range: 0.54–0.66). Within the WTA, the pairwise F_{ST} values indicated significant differentiation between sampling locations (mean $F_{ST}=0.01$, range: 0.005–0.013), with the U.S. Atlantic coast isolated from the northern Gulf of Mexico and Brazil (mean $F_{ST}=0.012$). In the IPA, all pairwise differentiation tests between Fiji (FIJ) and other sampling locations were significant (mean $F_{ST}=0.036$, range: 0.026–0.075), as well as comparisons between Iriomote Island (IRI) in Japan and other sampling locations (mean $F_{ST}=0.044$, range: 0.038–0.050). Among the other pairwise differentiations from the IPA, most were not significant with F_{ST} values indicating negligible genetic differentiation (mean $F_{ST}=0.003$, range: 0–0.011; Table 2), without clear geographic signal.

TABLE 2 Pairwise F_{ST} values. Dotted lines delimit the main biogeographic regions sampled in this study.

Sampling site	CRI	GOM	BRA	NAT	SAF	ARB	SEY	REU	LKA	TAI	DAR	PNG	CAP	SYD	IRI
GOM	0.357														
BRA	0.329	0.005													
NAT	0.392	0.010	0.013												
SAF	0.599	0.616	0.616	0.559											
ARB	0.627	0.631	0.626	0.572	0										
SEY	0.587	0.610	0.610	0.559	0	0.002									
REU	0.599	0.617	0.617	0.559	0.002	0	0.003								
LKA	0.656	0.646	0.638	0.591	0	0	0.001	0.003							
TAI	0.688	0.656	0.645	0.609	0	0	0.007	0.001	0.005						
DAR	0.619	0.625	0.621	0.563	0	0.002	0	0.003	0.006	0.004					
PNG	0.663	0.650	0.641	0.596	0	0.001	0.003	0.005	0	0.011	0.005				
CAP	0.605	0.622	0.621	0.566	0.005	0.004	0.007	0.005	0.003	0.006	0.005	0.006			
SYD	0.558	0.587	0.589	0.541	0	0	0	0.002	0	0.004	0	0	0.003		
IRI	0.602	0.623	0.623	0.577	0.047	0.046	0.040	0.042	0.049	0.050	0.044	0.044	0.040	0.038	
FJI	0.629	0.635	0.631	0.585	0.032	0.026	0.032	0.033	0.032	0.037	0.040	0.028	0.033	0.031	0.075

Note: Values significantly different from 0 after FDR correction ($\alpha=0.05$) are in bold.

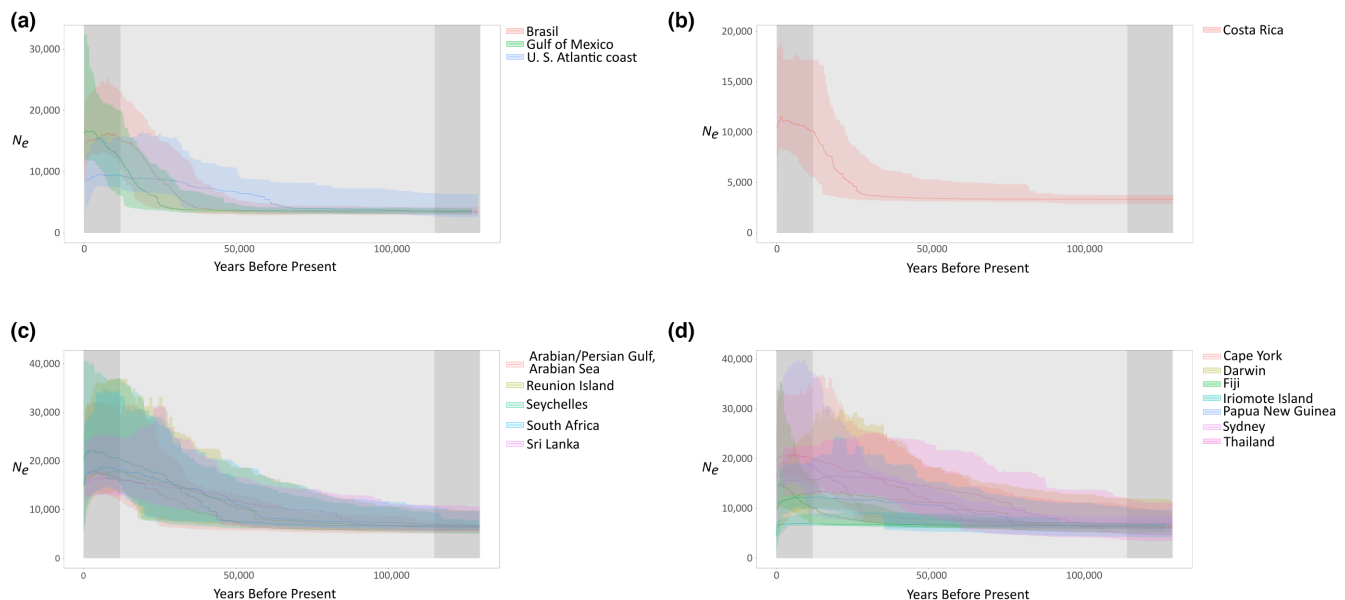


FIGURE 3 Variations in the median effective population size (N_e) through time and its 75% confidence interval estimated by STAIRWAYPLOT. The grey area indicates the last glacial period. (a) Western Tropical Atlantic; (b) Eastern Tropical Pacific; (c) Indo-West Pacific; and (d) Central Indo-Pacific.

3.3 | Demographic inferences

Summary statistics are presented in Table 1. TD were negative in all sampling locations, indicating an excess of low-frequency variants. According to the ABC-RF framework, the NS model had a posterior probability >0.5 in 10 sampling locations (one in the WTA and nine in the IPA; Table 1). The 95% credible intervals of T_{col} and N_{anc} for FIJ and IRI mostly overlapped prior distributions, suggesting that the data do not contain enough information to correctly estimate model parameters. Other IPA sampling locations showed N_e increase (mean N_{mod}/N_{anc} ratio \pm SE = 3.65 ± 0.48) occurring around $\sim 110,000$ B.P., switching from a mean N_{anc} of ~ 5000 individuals to a mean N_{mod} of $\sim 19,000$ individuals. The parameter estimation of the NS model for the WTA sampling site gave a similar pattern to most IPA sampling locations, with N_{anc} being approximately one-third of N_{mod} .

Demographic trajectories reconstructed with STAIRWAYPLOT (interpreted as N_e temporal variation because panmixia could not be rejected in most cases) were consistent with the ABC-RF results. The difference in the timing of the expansion stems from the fact that STAIRWAYPLOT implements a non-parametric N_e variation model, while the ABC-RF framework employs a single N_e time change. For all sampling locations except IRI, an increase in median N_e was observed, starting $\sim 20,000$ – $60,000$ B.P. in WTA populations (Figure 3a), while $\sim 60,000$ – $80,000$ B.P. in most of the IPA sampling locations (Figure 3c and Figure 3d), with Thailand being the oldest. Since then, a comparatively constant median N_e was observed (Figure 3) until a generalized reduction in recent generations. Three sampling locations from the IPA depart from this pattern: FIJ, IRI and Sydney (Figure 3d). FIJ and Sydney sampling locations fit the general template but with a younger expansion starting $\sim 20,000$ B.P.

IRI does not present any ancestral expansion, only N_e reduction in the last millennia.

3.4 | Simulation study—Detection of recent bottleneck

Coalescent simulations run under the NS model reproduced the genetic variability observed in real populations (Table 3, Data S3), and the simulated trajectory was generally well retrieved by the STAIRWAYPLOT (Data 5a), particularly when increasing the number of loci and sampled individuals. The STAIRWAYPLOT run on datasets simulated under the NS model presented a reduction in median N_e in the most recent (~ 10) generations when analysing 1000 to 5000 loci, as observed in real data (Figure 4a, Data 5a). This reduction disappeared when analysing more loci (Figure 4, Data S5). For scenarios simulated under NS_{BOT} , the STAIRWAYPLOT could recover a recent bottleneck ($T_{BOT}=5$) only for large N_e reduction ($BOT > 50\times$), showing a decreasing trajectory in recent generations (Figure 4b–e, Data 3b). In contrast, STAIRWAYPLOT reconstructed the decreasing N_e trajectory at all BOT intensities when datasets were simulated with older T_{BOT} . However, STAIRWAYPLOT progressively failed to recover the ancestral N_e expansion included in all scenarios as BOT and T_{BOT} values increased. For older ($T_{BOT}=450$ and 1500) and/or strongest bottlenecks ($BOT > 10\times$, Data 5d,e), the demographic history was dominated by the post-bottleneck coalescence rate: the STAIRWAYPLOT reconstructed populations with constant N_e corresponding to the post-bottleneck value (looking forward in time). For $T_{BOT}=1500$ and $BOT > 50\times$, almost all genetic diversity was lost and STAIRWAYPLOT could not reconstruct N_e trajectories over more than a few generations.

TABLE 3 Summary statistics averaged over 10 replicates of coalescent simulations of 5000 loci (100 bp each) under the NS and NS_{BOT} models with a mutation rate of 2.509×10^{-8} per site per generation.

T_{BOT}	BOT	TD	θ_π	θ_w	S
-	-	-0.89	6.9×10^{-4}	8.9×10^{-4}	1760.4
5	5	-0.86	7.0×10^{-4}	9.0×10^{-4}	1773.6
	10	-0.87	6.8×10^{-4}	8.8×10^{-4}	1734
	50	-0.80	6.8×10^{-4}	8.6×10^{-4}	1701.3
	100	-0.67	6.8×10^{-4}	8.2×10^{-4}	1623.4
50	5	-0.81	6.8×10^{-4}	8.5×10^{-4}	1688.2
	10	-0.69	6.8×10^{-4}	8.3×10^{-4}	1635.8
	50	-0.10	6.4×10^{-4}	6.6×10^{-4}	1306.6
	100	0.43	5.9×10^{-4}	5.3×10^{-4}	1056.1
450	5	-0.31	6.6×10^{-4}	7.2×10^{-4}	1426.8
	10	0.14	6.2×10^{-4}	6.0×10^{-4}	1183.6
	50	1.27	3.6×10^{-4}	2.7×10^{-4}	543
	100	1.32	1.9×10^{-4}	1.4×10^{-4}	282
1500	5	0.24	5.9×10^{-4}	5.6×10^{-4}	1109.6
	10	0.72	4.7×10^{-4}	3.9×10^{-4}	781.5
	50	0.86	1.0×10^{-4}	8.1×10^{-5}	160
	100	0.43	2.5×10^{-5}	2.2×10^{-5}	43.5

Abbreviations: BOT, reduction factor applied to N_{mod} ; S, number of segregating sites; T_{BOT} , onset of the bottleneck in number of generations; TD, Tajima's D; θ_w , Watterson's theta; θ_π , mean pairwise difference.

3.5 | Ancestral divergence

Model selection identified the scenario of an ancestral divergence of IPA as the most likely (Figure 2 – Model 2, Data S4). Within this topology, we found that the model with highest support displayed continuous asymmetrical migration rates between genetic clusters (Table 4). According to this model (Table 4), the estimated divergence time between the WTA and EPA was ~40,000 B.P., and ~56,000 B.P. between IPA and the ancestor of the EPA and WTA genetic clusters. IPA estimated modern N_e closer to other demographic inferences performed in this study (~14,500 individuals) to the contrary of the other two clusters which presented lower N_e estimates (EPA = ~7500 and WTA = ~3500). Similarly, ancestral N_e were small, below 300 individuals in both cases. Migration rates were extremely low, less than one individual per generation in all cases. IPA estimated modern N_e falls outside its 95% bootstrap confidence interval, as did the estimated divergence time between IPA and the ancestor of EPA and WTA. This indicated a lack of power to infer these parameter values with confidence.

4 | DISCUSSION

Previous studies have used microsatellites, mtDNA or genomic markers to uncover the mechanisms driving gene flow in *C. leucas* (Devloo-Delva et al., 2023; Glaus et al., 2020; Pirog, Jaquemet, et al., 2019;

Pirog, Ravigné, et al., 2019; Testerman, 2014). These studies underlined the weak and/or non-significant genetic differentiation between sampling locations inside biogeographic regions while suggesting a strong disjunction among them. However, reconstruction accuracy increases with the number of independent loci analysed (Felsenstein, 2006; Nordborg, 2019; Wakeley, 2009) and a representative sampling across the biogeographic range, which helps refine our understanding of the evolutionary history of the Bull Shark.

4.1 | *Carcharhinus leucas* biogeography

The present study supports the importance of biogeographic barriers in the diversification of *C. leucas* (Devloo-Delva et al., 2023). According to our estimations, the divergence of the IPA from the WTA and EPA occurred at ~55,000 B.P., while the divergence between the WTA and EPA occurred ~40,000 B.P. It is worth noting that Pirog, Jaquemet, et al. (2019); Pirog, Ravigné, et al. (2019) timed the divergence of IPA and WTA at ~1.23 million B.P. using mtDNA, linking it to the formation of the Benguela Current. The difference in the estimated divergence dates with the known insurgence of biogeographic barriers, that is, the Isthmus of Panama, the Benguela Current and the Eastern Pacific open ocean, is difficult to reconcile. While the Benguela Current is a permeable barrier (Bernard et al., 2018; Lesturgie, Lainé, et al., 2022; Reid et al., 2016), the divergence between EPA and WTA after the closure of the Isthmus of Panama is surprising (but see Galván-Quesada et al., 2016). Indeed, low water temperatures form a thermal barrier to *C. leucas* around the southern and northern tips of the American continent for millions of years. Two scenarios could explain this discrepancy. The first is a secondary contact between EPA and WTA, artificially decreasing their divergence time via genomic introgression. To test this hypothesis, secondary contact scenarios were added to *fastsimcoal* modelling under the hypothesis that the opening of the Panama Canal fuelled migration between ocean basins. However, the likelihood of these models was lower (Data S4) and the estimated divergence was still too recent (data not shown). A second explanation is that the mutation rate used here was two orders of magnitude higher than the real one. It seems unlikely that *C. leucas* would have such a slow mutation rate (around 10^{-10} per site per generation), which would be the lowest documented so far in vertebrates. Ultimately, our set of loci represented a fraction of the nuclear genome and in some cases the parameter estimates fell outside the confidence intervals, therefore the estimates should be taken cautiously. However, this does not affect the model selection procedure. Whole-genome sequencing will certainly help refine our estimates.

4.2 | Population structuring

This study further confirms that *C. leucas* is divided into at least three stocks (Olver et al., 1995) corresponding to major marine biogeographic regions: WTA, EPA and IPA. Indeed, all analyses showed that while there is high gene flow within regions, they are almost completely

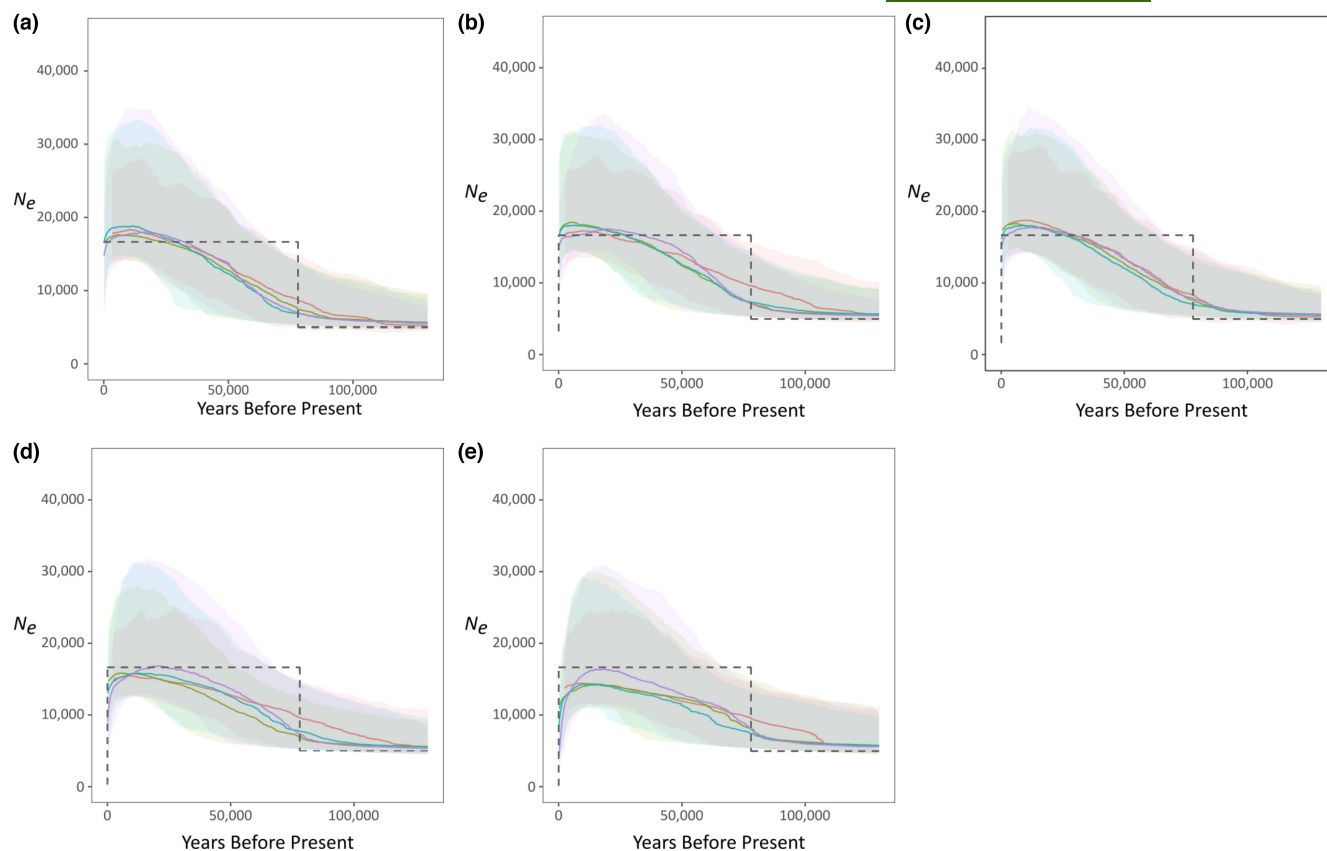


FIGURE 4 STAIRWAYPLOT estimates averaged over 10 replicates of the effective population size (N_e) variations through time of scenarios: (a) NS; (b) NS_{BOT} with a bottleneck starting five generations ago with intensity $BOT = 5$; (c) $BOT = 10$; (d) $BOT = 50$; and (e) $BOT = 100$. The median values are presented in bold and their 75% confidence intervals as shaded areas. All scenarios are based on coalescent simulations of 5000 loci (100 bp each) with a mutation rate of 2.509×10^{-8} per site per generation of 5 (red), 10 (green), 15 (blue) and 20 (purple) diploid individuals. The grey dotted line represents the true (simulated) N_e variation through time.

TABLE 4 Parameter estimation using *fastsimcoal* under the most likely divergence model.

Parameters	Initial boundaries	Estimated value [95% confidence interval]
N_{e1} : WTA modern N_e	0.5–50,000	3486 [1637–4397]
N_{e2} : EPA modern N_e	0.5–50,000	7284.5 [5020–7179]
N_{e3} : IPA modern N_e	0.5–50,000	14,484 [9928–10,681]
T_{12} : divergence time between WTA and EPA	130–1300,000	40,937 [36,626–76,837]
N_{anc2} : WTA-EPA ancestral N_e	0.5–50,000	135 [118–1726]
T_{23} : divergence between WTA-EPA and IPA	130–1300,000	56,121 [173,725–261,880]
N_{anc1} : WTA-EPA-IPA ancestral N_e	0.5–50,000	218 [126–2997]
Migration rate from WTA to EPA	$1 \times 10^{-7} - 0.01$	0.31 [0.0020–1.61]
Migration rate from EPA to WTA	$1 \times 10^{-7} - 0.01$	0.026 [0.013–0.50]
Migration rate from WTA to IPA	$1 \times 10^{-7} - 0.01$	0.13 [0.042–0.36]
Migration rate from IPA to WTA	$1 \times 10^{-7} - 0.01$	0.019 [0.0014–0.011]
Migration rate from IPA to EPA	$1 \times 10^{-7} - 0.01$	0.0032 [0.0018–0.015]
Migration rate from EPA to IPA	$1 \times 10^{-7} - 0.01$	0.059 [0.034–0.11]

Note: Composite maximum-likelihood estimates of effective population sizes are presented in number of (diploid) individuals per population, divergence times in number of years (generation time = 13 years) and migration rates are expressed as number of migrants per generation (Nm , backward in time). One hundred parametric bootstrap replicates were used to calculate the 95% confidence intervals.

Abbreviations: EPA, Eastern Tropical Pacific; IPA, Western and Central Indo-Pacific; WTA, Western Tropical Atlantic.

genetically isolated. The lack of genetic differentiation inside regions is probably related to *C. leucas* ecology, as it is capable of moving thousands of kilometres along continents (Espinoza et al., 2016, 2021) and in the open ocean (Lea et al., 2015). However, two IPA locations stood out: FIJ and IRI. Devloo-Delva et al. (2023) and Glaus et al. (2020) identified FIJ as genetically distinct from other IPA locations, the latter suggesting that it resulted from the archipelago's oceanic isolation. However, FIJ was not the only isolated sampling location; Seychelles is ~1000km apart from Madagascar, yet it did not show traces of genetic isolation. Sampling bias could explain FIJ genetic differentiation, as most samples used here came from intermittently resident females suspected to pup in the area (Bouveroux et al., 2021; Brunnschweiler & Barnett, 2013; Cardenosa et al., 2017; Glaus et al., 2019). Given the suspected reproductive philopatric behaviour of *C. leucas* females (Devloo-Delva et al., 2023; Espinoza et al., 2016; Pirog, Jaquemet, et al., 2019; Pirog, Ravigné, et al., 2019), the Fijian genetic distinctiveness could stem from relatedness, as in Lemon sharks (Feldheim et al., 2014). Likewise, samples from Iriomote Island originate from a river used as nursery, with most individuals sampled younger than 2 years old (data not shown). However, if the genetic isolation came from higher relatedness, we would expect strong and positive F_{IS} values, which were not observed anywhere (Data S1). Ultimately, we could not exclude the presence of undiscovered biogeographic barriers, or that the high differentiation of populations at the edge of the species distribution is due to a recent range expansion. Sampling the northern IPA and Micronesia would determine whether the pattern corresponds to actual biogeographical borders or sampling artefacts (Gausmann, 2021). The genetic differentiation between the U.S. Atlantic coast and other WTA sampling locations could result from a biogeographic border, as the Florida Peninsula forms a barrier between the Gulf of Mexico and the Atlantic (Hirschfeld et al., 2021). A temporal effect could also play part in the distinction of this population (some samples were collected between 1984 and 1987).

Sharks life history traits largely affect population structure (Lesturgie, Lainé, et al., 2022; Lesturgie, Planes, & Mona, 2022), but few species have been studied on a global scale with genomic datasets similar to the one presented here. Species exhibiting strict fidelity to coral reefs such as the Blacktip Reef and Grey Reef sharks present strong genetic structuring over the Indo-Pacific (Lesturgie et al., 2023; Maisano Delser et al., 2019). On the contrary, Tiger Shark, presenting a similar circumtropical distribution to *C. leucas*, is divided into two almost independent panmictic stocks (the Atlantic and the Indo-Pacific, Lesturgie, Planes, & Mona, 2022; Lesturgie, Lainé, et al., 2022). More species with similar distribution and such extensive geographic coverage need to be studied to better understand the relationship between life history traits and genetic structuring.

4.3 | Demographic history and effective population size

C. leucas global N_e increased during the last glacial period. During this period, sea levels were at least 50m below present,

extending coastlines and so the available habitat for *C. leucas* (Carlson et al., 2010; Graham et al., 2016; Hammerschlag et al., 2012; Heupel et al., 2015; Niella et al., 2020), potentially supporting larger populations. As *C. leucas* inhabits areas with water temperatures down to 18°C (Brunnschweiler et al., 2010; Lea et al., 2015; Matich & Heithaus, 2012; Smoothey et al., 2016, 2019, 2023), sea surface temperature changes during the LGM did not significantly reduce its distribution in the tropics (Monteagudo et al., 2021). Additionally, long-range movements (Espinoza et al., 2016; Lea et al., 2015; Lee et al., 2019) may have facilitated colonization of newly emerged areas. However, if available habitat was the sole driver of N_e , a reduction should have been observed after the LGM, as available coastal habitats receded. The ability to detect bottlenecks depends on many factors: intensity, timing, ancestral demography and the number of individuals and loci sampled. In addition, recent population declines are harder to detect for long-lived and late-maturing species, as fewer generations have elapsed in the same amount of time (e.g. Roman & Palumbi, 2003). Our simulations under the NS_{BOT} model suggested that even a limited N_e reduction starting during the mid-Holocene or the LGM (Data S5d,e) would have hidden the ancestral expansion retrieved in our populations, and it is therefore inconsistent with *C. leucas* evolutionary history. Conversely, our dataset does not have enough power to detect recent N_e reduction, at least with the use of the folded SFS. Indeed, the small decrease observed in the recent generations is most likely an artefact (Data S5a). In the future, it will be important to complement SFS-based methods with those based on linkage disequilibrium statistics, better suited to detect recent changes in N_e (Boitard et al., 2016; Kerdoncuff et al., 2020; Santiago et al., 2020) and to develop full-genome resources.

4.4 | Perspective of *C. leucas* populations conservation and management

Based on this study and complementing previous findings (Devloo-Delva et al., 2023; Pirog, Jaquemet, et al., 2019; Pirog, Ravigné, et al., 2019), *C. leucas* from the IPA, WTA and EPA form three independent genetic clusters, and should be considered as independent stocks following Olver et al. (1995). Demographic modelling showed that the species still harbours significant genetic diversity, globally retaining its evolutionary potential, according to Frankham et al. (2014). Interestingly, the IPA seemed to be the oldest cluster, harbouring the highest genetic diversity and likely being the centre of origin of this species. Two important results are highlighted by our simulations: (i) decrease in N_e after the LGM or mid-Holocene can be excluded, as it would have shown a detectable signature on the observed genetic variation; (ii) conversely, a bottleneck starting five generations ago is undetectable with a dataset of this size, unless its strength approaches extreme values (Figure 4, Data S5b). The ongoing population depletion in the IPA may not be recovered using the panel of loci analysed here. In conclusion, even though elasmobranch populations have been following a downward trend for several decades (Dulvy

et al., 2021; Pacoureau et al., 2021), its impact on the genetic diversity of this species requires more genomic data and the application of linkage disequilibrium-based statistics to be detectable.

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CONFLICT OF INTEREST STATEMENT

The authors declare no conflict of interest.

DATA AVAILABILITY STATEMENT

The original 513 FASTQ files (475 individuals) and their metadata, including the final list of 309 individuals used to generate the results of this study, are available on DataDryad (<https://doi.org/10.5061/dryad.9zw3r22mn>).

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BIOSKETCH

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conceived the ideas; all authors collected the data; BDP, FDD, PL and SM analysed the data; and BDP, FDD, HM, PF and SM led the writing.

SUPPORTING INFORMATION

Additional supporting information can be found online in the Supporting Information section at the end of this article.

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